<< <u>2023-01-19</u> | <u>2023-01-21</u> >>

Decision is a risk rooted in the courage of being free. — Paul Tillich

Lectures:

<u>Simulation and modeling of natural processes | 自然过程的仿真与建模 |</u>
 <u>Python | ABM 哔哩哔哩 bilibili</u>

Chapter 1 Basic Simulation Modeling

323_pagefiles1C H A P T E R 1 Basic Simulation Modeling Recommended sections for a fi rs <u>show annotation</u>

1.1 The Nature of Simulation

, 1.81.1THE NATURE OF SIMULATION This is a book about techniques for using computers to imitate, or simulate, the operations of various kinds of real-world facilities or processes. The facility or pro-cess of int <u>show annotation</u>

of interest; this is called an <mark>analytic solu-tion</mark> . However, most real-world syste

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解析解: exact information on questions of interest

files2 basic simulation modeling <mark>Application areas for simulation are numerous and diverse</mark> . Below is a list of some partic <u>show annotation</u>

Application areas for simulation:

- Designing and analyzing manufacturing systems (制造系统的设计与分析);
- Evaluating military weapons systems or their logistics requirements (军用武 器系统及其后勤保障的评估);
- Determining hardware requirements or protocols for communications networks (通信网络中硬件需求或协议的确定);
- Determining hardware and software requirements for a computer system (计算机系统软硬件需求的确定);
- Designing and operating transportation systems such as airports, freeways, ports, and subways (运输系统(如机场、高速公路、港口、地铁等)的设计 与运营):
- Evaluating designs for service organizations such as call centers, fast-food restaurants, hospitals, and post offices (服务组织(如呼叫中心、快餐店、医 院、邮局等)的评估设计);
- Reengineering of business processes (业务流程的重组);
- Analyzing supply chains (供应链分析);
- Determining ordering policies for an inventory system (库存系统订货策略的 确定);
- Analyzing mining operations (采矿作业的分析)。

an analytic solu-tion. However, most real-world systems are too complex to allow realistic models to be evaluated analytically, and these models must be studied by means of simulation. In a simulation we use a <u>show annotation</u>

oes not work out. However, a <mark>careful simulation study could shed some light on the question</mark> by simulating the operation <u>show annotation</u>

s, if not the most widely used. One indication of this is the Winter Simulation Conference, which attracts 600 to 800 people every year. In addition, there are sever <u>show annotation</u>

d.There have been, however, several impediments to even wider acceptance and usefulness of simulation . First, models used to study la <u>show annotation</u>

缺点

缺点:

- 用于大规模系统的模型变得非常复杂
- 仿真往往需要耗费大量时间。但随着计算机运行速度的不断加快,费用逐渐降低,逐渐迎难而解。

and usefulness of simulation. First, models used to study large-scale systems tend to be very complex , and writing computer progr <u>show annotation</u>

program" a simulation model. <mark>A second problem with simulation of complex systems is that a large amount of computer time is some-times required. However, this diffi culty h <u>show annotation</u></mark>

become faster and cheaper. Finally, there appears to be an unfortunate impression that simulation is just an exercise in computer programming, albeit a complicated one . Consequently, many simulation <u>show annotation</u>

he latter chapters of this book. Perspectives on the historical evolution of simulation modeling may be found in Nance and Sargent (2002). In the remainder of this chapter <u>show annotation</u>

rying degrees of complexity. <mark>All of the computer code shown in this chapter can be downloaded from <u>www.mhhe.com/law</u>. 1.2SYSTEMS, MODELS, AND SIMULATI <u>show annotation</u></mark>

code url

#code

1.2 Systems, Models, and Simulation

系统分类:

- 离散系统: 状态变量在一些离散时间点上瞬时变化的系统;
- 连续系统: 状态变量随时间变化而变化的系统。

of each cus-tomer in the bank. <mark>We categorize systems to be of two types, discrete and continuous.</mark> A discrete system is one for <u>show annotation</u>

types, discrete and continuous. <mark>A discrete system is one for which the state variables change instantaneously at separated points in time</mark>. A bank is an example of a disc <u>show annotation</u>

ishes being served and departs. <mark>A continuous system is one for which the state variables change continuously with respect to time.</mark> An airplane moving through th show annotation

calsolution SimulationFIGURE 1.1 Ways to study a system. Law01323_ch01_001-084.indd Page show annotation

Ways to study a system

- 用*实际系统*的实验与用*系统模型*的实验 (Experiment with the Actual System vs. Experiment with a Model of the System)
- 物理模型与数学模型 (Physical Model vs. Mathematical Model).
- 解析解与仿真 (Analytical Solution vs. Simulation).

t is useful for this purpose to <mark>classify simulation models along three different dimensions:</mark> • Static vs. Dynamic Simulation <u>show annotation</u>

Classify simulation models along three different dimensions:

- 静态与动态仿真模型 (Static vs. Dynamic Simulation Models)
- 确定的与随机的仿真模型 (Deterministic vs. Stochastic Simulation Models)
- *连续的*与离散仿真模型 (Continuous vs. Discrete Simulation Models).

The disadvantage of simulation:

ems are modeled stochastically. Stochastic simulation models produce output that is itself random, and must therefore be treated as only an estimate of the true characteristics of the model; this is one of the main disadva show annotation

1.3 Discrete Event Simulation

ty.)1.3DISCRETE-EVENT SIMULATION Discrete-event simulation concerns the modeling of a system as it evolves over time by a representation in which the state variables change instantaneously at separate points in time. (In more mathematical terms show annotation

time from one value to another. We call the variable in a simulation model that gives the current value of simulated time the simulation clock . The unit of time for the simul show annotation

仿真钟,仿真模型中给出仿真时间当前值的这个变量

仿真钟: 仿真模型中给出仿真时间当前值的这个变量

on the computer.Historically, <mark>two principal approaches have been suggested for advancing the simulation clock: next-event time advance and fi xedincrement time advance.</mark> Since the fi rst approach is us

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两种方式推进仿真钟:下一最早发生事件时间推进和固定增量时间推进

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omputer program. In particular, the following components will be found in most discrete-event simula-tion models using the next-event time-advan show annotation

离散事件仿真模型的组件:

- 系统秋态: 播述在某一特定时刻系统所必需的一组状态变量的集合。
- 伤真钟:给出仿真时间当前值的变量。
- 事件表: 包含下一次每类事件发生时间的表。
- 统计计数器:用于存储系统性能统计信息的变量。
- 初始化例程:零时刻对伤真模型进行初始化的子程序。
- 定时例程:用于从事件表中确定下一事件并将伤真钟推进到该下一事件发生时 间的一个子程序。
- 事件例程;当某一特定类型事件发生时更新系统状态的子程序(每类事件有一 个事作例程)。
- 库例程:根据确定的概率分布产生随机观测值的一组子程序,作为仿真模型的 部分。
- 报告生成器:计算性能期望度量的估计值(通过统计计数器),并在仿真结束
 时生成一个报告的一个子程序。
- 主程序:调用定时例程来确定下一事件,然后将控制转移到相应的事件例程以
 适当更新系统状态的一个子程序。主程序还可以检查终止条件并在终止时调用
 报告生成器。

1.4 Simulation of a Single-server Queueing System

d service-time random variables. To measure the performance of this system, we will look at estimates of three quantities. First, we will estimate the expe show annotation

To measure the performance of the discrete simulation system:

- $\underline{\psi}$ $\underline{\psi}$
- *队列中顾客平均数的期望值*: the expected average number of customers in the queue (but not being served), denoted by q(n).
- *服务台的期望利用率*: The expected utilization of the server is the expected proportion of time during the simulation.

estimates of three quantities. First, we will estimate the expected average delay in queue of the n cus-tomers completing their delays during the simulation; we denote this quantity by d(n). The word "expected" in the defi show annotation

ure for our simple model here is the expected average number of customers in the queue (but not being served), denoted by q(n), where the n is necessary in the <u>show annotation</u>

sure of how busy the server is. The expected utilization of the server is the expected pro-portion of time during the simulation [from time 0 to time T(n)] that the server is busy (i.e., show annotation

in a data-processing operation. To recap, the three measures of performance are the average delay in queue d[^](n), the time-average number of customers in queue q[^] (n), and the proportion of time the server is busy u[^] (n). The average delay in queue is a <u>show annotation</u> Measurements

#Measurements

1.5 Simulation of An Inventory System

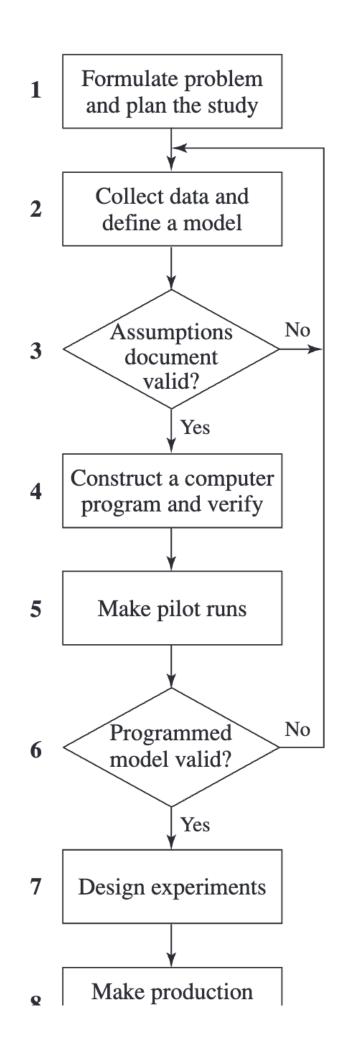
1.6 Parallel/Distributed Simulation and The High Level Architecture

iles62 basic simulation modeling <mark>logic is executed in order of the events' simulated time of occurrence; i.e., the simu-lation is sequential.</mark> Furthermore, all work is done o <u>show annotation</u>

sequential. a. 串行的

1.7 Steps in a Sound Simulation Study

ter one 67 sequential process. <mark>As one proceeds with the study, it may be necessary to go back to a previous step.</mark> 1. Formulate the problem and <u>show annotation</u>



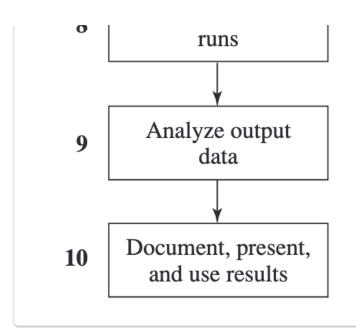


FIGURE 1.46 Steps in a simulation study.



- 1. 问题形式化与制定研究计划
- 2. 收集数据并定义模型
- 3. 假设文档是有效的吗
- 4. 编制计算机程序并校验
- 5. 进行操控运行
- 6. 编程模型有效吗
- 7. 设计实验
- 8. 进行批次运行
- 9. 分析输出数据
- 10. 形成文件, 提交及使用结果

1.8 Advantages, Disadvantages and Pitfalls of Simulation

od for studying complex systems. Some possible advantages of simulation that may account for its widespread appeal are the following. • Most complex, real-world syst <u>show annotation</u>

ead appeal are the following.• Most complex, real-world systems with stochastic elements cannot be accurately described by a mathematical

model that can be evaluated analytically . Thus, a simulation is often t <u>show annotation</u>

无法求出解析解

Law01323_pagefileschapter one 71 Simulation is not without its drawbacks. Some disadvantages are as follows .• Each run of a stochastic sim <u>show annotation</u>

2. Chapter 2 Modeling Complex Systems

han a random-number generator). Most real-world systems, however, are quite complex, and coding them without supporting software can be a diffi cult and time-consuming task .In this chapter we first discu <u>show annotation</u>

2.2 List Processing in Simulation

e.In Sec. 2.2.1 we discuss <mark>two approaches to storing lists of records in a computer—sequential and linked allocation</mark> —and then explain why the l <u>show annotation</u>

存储记录表的两种方式:

- 顺序分配 (Sequential allocation)
- 链式分配 (Linked allocation)

3. Simulation Software

3.1 Introduction

2, the reader probably noticed several features needed in programming most discrete-event simulation models, including: • Generating random numbers, <u>show annotation</u> several features needed in programming most discrete-event simulation models, including:

- Generating random numbers, that is, observations from a U(0,1)probability distribution (产生随机数,即均匀概率分布 U(0,1) 的观测值);
- Generating random variates from a specified probability distribution (e.g., exponential) (产生一个特定概率分布(例如,指数分布)的随机变量);
- Advancing simulated time (推进仿真时间);
- Determining the next event from the event list and passing control to the appropriate block of code (从事件表中确定下一事件,并将控制权转交给适当 的代码块);
- Adding records to, or deleting records from a list (向一个表添加记录或从表 中删除记录);
- Collecting output statistics and reporting the results (搜集输出统计信息并生成结果报告);
- Detecting error conditions (探测错误发生的条件)。

, are described in Sec. 3.4. Section 3.5 gives brief descriptions of Arena, ExtendSim, and Simio, which are popular general-purpose simulation packages . A simulation model of a small show annotation

3.3 Classification of Simulation Software

ion-Oriented Simulation Packages There are two main types of simulation packages for discrete-event simulation, namely, general-purpose simulation software and application-oriented simulation software. A general-purpose simulation pa <u>show annotation</u>

grams in Chaps. 1 and 2, w <mark>e used the event-scheduling approach to</mark> discrete-event simulation modeling . A system is modeled by identif show annotation

事件调度

4. Review of Basic Probability and Statistics

4.1 Introduction

roposed system confi guration. The use of probability and statistics is such an integral part of a simulation study that every simulation modeling <u>show annotation</u>

experiments (Chaps. 11 and 12). In this chapter we establish statistical notation used throughout the book and review some basic probability and statistics particularly relevant to simulation. We also point out the potential <u>show annotation</u>

4.2 Random Variables and Their Properties

Notations:

- Sample space S: The set of all possible outcomes of an experiment
- Random variable: a function (or rule) that assigns a real number (any number greater than $-\infty$ and less than ∞) to each point in the sample space S

In general, we denote random variables by capital letters such as X, Y, Z and the values that random variables take on by lowercase letters such as x, y, z.

- Distribution function (sometimes called the cumulative distribution function) F(x).
- Joint probability mass function p(x, y)
- Joint probability density function (f(x, y))
- (Marginal) Probability density function $f_X(x)$

- Mean or Expected value of X_i is μ_i or $\mathbb{E}(X_i)$
- Variance σ_i^2 or $Var(X_i)$

lue no larger than the number x. <mark>A distribution function F(x) has the following properties</mark> :1. 0 # F(x) # 1 for all x.2. <u>show annotation</u>

Properties:

- $0 \leq F(x) \leq 1, orall x$
- If $x_1 < x_2$, then $F(x_1) \leq F(x_2)$
- $\lim_{x o \infty} F(x) = 1$ and $\lim_{x o 0} F(x) = 0$

X 5 x, Y 5 y) for all x, ywhere p(x, y) is called the joint probability mass function of X and Y. In this case, X and show annotation

e function f (x, y), called the <mark>joint probability density function</mark> of X and Y, such that for all s <u>show annotation</u>

fY (y) 5 # 2 f(x, y) dxare the (marginal) probability density functions of X and Y, respectively. E X A M P L E 4show annotation

$$egin{aligned} f_X(x) &= \int_{-\infty}^\infty f(x,y) \ \mathrm{d} y \ f_Y(y) &= \int_{-\infty}^\infty f(x,y) \ \mathrm{d} x \end{aligned}$$

two random variables Xi and Xj. The mean or expected value of the random variable Xi (where i 5 1, 2, . . . , n) will <u>show annotation</u>

Then the following are important properties of means:

1.
$$\mathbb{E}(cX) = c\mathbb{E}(X)$$
.

2. $\mathbb{E}(\sum_{i=1}^n c_i X_i) = \sum_{i=1}^n c_i \mathbb{E}(X_i)$ even if the X_i 's are dependent (相关的).

note a constant (real number). T <mark>hen the following are important properties of means</mark> :1. E(cX) 5 cE(X).2. E(Oni 5 1

<u>show annotation</u>

e unique for some distributions. The variance of the random variable Xi will be denoted by s2i or Var(Xi) and is defi ned by s2i 5 E[(Xi show annotation

5 E(X2) 2 m2 5 13 2 a 12 b25 112 The variance has the following properties: 1. Var(X) \$ 0.2. Var(cX) 5 c2 <u>show annotation</u>

Properties of variance:

• $\operatorname{Var}(X_i) \geq 0$

•
$$\operatorname{Var}(cX_i) = c^2 \operatorname{Var}(X_i)$$

• $\operatorname{Var}(\sum_{i=1}^{n} X_i) = \sum_{i=1}^{n} \operatorname{Var}(X_i)$ if the Xi's are independent (or uncorrelated, as discussed below).

In particular, suppose that Xi has a normal distribution with mean μ_i and standard deviation σ_i . In this case, for example, the probability that X_i is between $\mu_i - 1.96\sigma_i$ and $\mu_i + 1.96\sigma_i$ is 0.95.

distribution (see Sec. 6.2.2). In particular, suppose that Xi has a normal distribution with mean mi and standard deviation si. In this case, for example, the prob-ability that Xi is between mi 2 1.96si and mi 1 1.96si is 0.95 .We now consider measures of dep <u>show annotation</u>

between two random variables. T <mark>he covariance between the random variables Xi and Xj (where i 5 1, 2, . . . , n; j 5 1, 2, . . . , n)</mark> , which is a measure of their (I <u>show annotation</u>

Covariance:

$$C_{ij} = \mathbb{E}\left[(X_i - \mu_i)(X_j - \mu_j)
ight] = \mathbb{E}(X_iX_j) - \mu_i\mu_j$$

Note: covariances are symmetric, that is, $C_{ij} = C_{ji}$, and if i = j, then $C_{ij} = C_{ji} = \sigma_i^2$.

5 215 2 a 25 b a 25 b 5 2 275 If Cij = 0, the random variables Xi and Xj are said to be uncorrelated. It is easy to show that if X show annotation

units of minutes squared.) As a result, we use the correlation rij, defi ned b y rij 5 Cij2s2i s2j i 5 1, 2, . show annotation

Correlation ρ_{ij} or $Cor(X_i, X_j)$:

$$ho_{ij} = rac{C_{ij}}{\sqrt{\sigma_i^2 \sigma_j^2}}$$

4.3 Simulation Output Data and Stochastic Process

ata analysis in Chaps. 9 and 10. A stochastic process is a collection of "similar" random variables ordered over time, which are all defi ned on a common sample space. The set of all possible val-ues that these random variables can take on is called the state space . If the collec-tion is X1, X2, <u>show annotation</u>

Note: 随机过程是按时间排序的"类似的"随机变量的几何,随机变量全部定义在共同 样本空间上。这些随机变量可以取的所有可能值的集合称为状态空间。

Example: For a single-server queueing system, e.g., M/M/1 queue, with IID inter-arrival times A_1, A_2, \dots , IID service times S_1, S_2, \dots and customers served in a FIFO manner. 定义在队列中的延误时间 (delays in queue) D_1, D_2, \dots as follows:

 $D_{1} = 0$

$$D_{i+1} = \max\{D_i + S_i - A_{i+1}, 0\} \hspace{0.4cm} ext{for} i = 1, 2, \cdots$$

Note: D_i and D_{i+1} are positively correlated.

- 令 *Q*(*t*) 为在 *t* 时刻在队列中的顾客数,则 {*Q*(*t*), *t* ≥ 0} 是一个连续时间随机
 过程,其状态空间为 {0,1,2,···}
- 令 *C_i* 为第 *i* 个月的总成本(如订货、存储和缺货成本的总和),则 *C*₁,*C*₂,…
 是一个离散时间随机过程,其状态空间为非负实数。

ce the nonnegative real numbers. To draw inferences about an underlying stochastic process from a set of simula-tion output data, one must sometimes make assumptions about the stochastic proces s that may not be strictly true

show annotation

需要预设随机过程的假设

property that we now defi ne <mark>. A discrete-time stochastic process X1, X2, . . . is said to be covariance-stationary</mark> if mi 5 m for i 5 1, 2, . .

show annotation

o

一个常见的例子是假设随机过程是协方差平稳的。Definition:

 $\mu_i=\mu \quad ext{for } i=1,2,\cdots ext{ and } -\infty <\mu <\infty$

$$\sigma_i^2=\sigma^2 \quad ext{for } i=1,2,\cdots ext{ and } \sigma^2<\infty$$

因此,对于一个协方差平稳过程,均值和方差是不随时间变化而变化的,且两个观 测值 X_i 和 X_{i+j} 之间的协方差仅仅取决于间隔 j,而不取决于实际时间值 i和 i+j

t), t \$ 0} in an analogous way.) For a covariance-stationary process, we denote the covariance and co show annotation

For a covariance-stationary process, we denote the covariance and correlation between X_i and X_{i+j} by C_j and ρ_j , respectively, where

$$ho_j = rac{C_{i,i+j}}{\sqrt{\sigma_i^2 \sigma_{i+j}^2}} = rac{C_j}{\sigma^2} = rac{C_j}{C_0} \quad ext{ for } j = 0, 1, 2, \dots$$

4.4 Estimation of Means, Variances, and Correlations

lation variance s2 and that our primary objec-tive is to estimate m; the estimation of s2 is of seco show annotation

For IID random variables

• Sample mean

$$ar{X}(n) = rac{{\sum\limits_{i = 1}^n {X_i } }}{n}$$

• Sample variance

$$S^2(n)=rac{\displaystyle\sum\limits_{i=1}^n \left[X_i-ar{X}(n)
ight]^2}{n-1}$$

• Variance of sample mean

$$\mathrm{Var}[ar{X}(n)] = rac{\sigma^2}{n}$$

• Unbiased estimator of $Var[\bar{X}(n)]$

$$\widehat{\operatorname{Var}}[ar{X}(n)] = rac{S^2(n)}{n} = rac{\sum_{i=1}^n \left[X_i - ar{X}(n)
ight]^2}{n(n-1)}$$

For covariance-stationary stochastic process

Sample variance

$$\mathrm{E}[S^2(n)] = \sigma^2 \left[1 - 2 rac{\sum\limits_{j=1}^{n-1} (1-j/n)
ho_j}{n-1}
ight]$$

• Variance of the sample mean

$$\mathrm{Var}[ar{X}(n)] = \sigma^2 rac{\left[1+2\sum\limits_{j=1}^{n-1}(1-j/n)
ho_j
ight]}{n}$$

31X1, X2, . . . , Xn are from a covariance-stationary stochastic process . Then it is still true that the <u>show annotation</u>

s often been done historically, <mark>there are two sources of error: the bias in S</mark> 2(n) as an estimator of s2 and the negligence of the correlation terms in Eq. (4.7). As a matter of fa <u>show annotation</u>

4.5 Confidence Intervals and Hypothesis Tests for the Mean

and s2 5 1 (henceforth called a <mark>standard normal random variable;</mark> see Sec. 6.2.2), is given by £(<u>show annotation</u>

• Standard normal random variable:

$$\Phi(z) = rac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{rac{-y^2}{2}} \,\mathrm{d}\, y \qquad -\infty < z < \infty$$

re larger.) Notice also that for <mark>a particular n, coverage decreases as the</mark> skewness of the distribution gets larger , where skewness is defi ned by show annotation

• skewness

$$u = rac{E\left[(X-\mu)^3
ight]}{\left(\sigma^2
ight)^{3/2}} \quad -\infty <
u < \infty$$

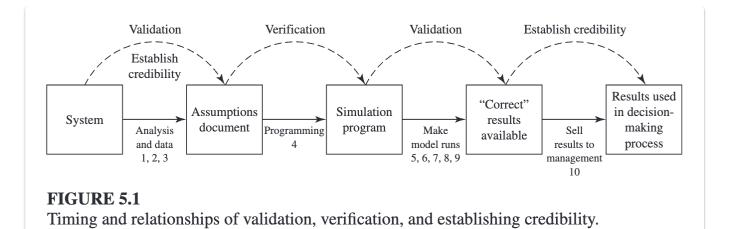
5. Building Valid, Credible, and Appropriately Detailed Simulation Models

5.1 Introduction and Definitions

5.1 INTRODUCTION AND DEFINITIONS One of the most diffi cult problems facing a simulation analyst is that of trying to determine whether a simulation model is an accurate representation of the actual system being studied, i.e., whether the model is valid . In this chapter we present a p show annotation

d Knepell and Arangno (1993). We begin by defining the important terms used in this chapter, including verifica-tion, validation, and credibility. Verifi cation is concerned with show annotation

- Verification is concerned with determining whether the "assumptions document"
- Validation is the process of determining whether a simulation model is an accurate representation of the system, for the particular objectives of the study.
- A simulation model and its results have credibility if the manager and other key project personnel accept them as "correct."



5.2 Guidelines for Determining the Level of Model Detail

ve computer memory requirements. We now present some general guidelines for determining the level of detail required by a simulation model [see also Law (1991) and R <u>show annotation</u>

确定仿真模型所需详细程度的某些一般准则:

- 认真定义要研究的问题的细节以及将用于评估的性能度量;
- 实体通过放着呢模型并不总是必须于是提通过对应系统相同;
- 使用主题专家 (SME) 和灵敏度分析来帮助决定模型的详细程度;
- 建模初学者常犯的错误是模型的细节过多;
- 在模型中不要有比解决关注的问题所必须的更多细节,前提是模型必须足够详 细保证其是可信的;
- 模型的详细程度应该与可用数据的类型相一致;
- 在实际的所有仿真研究中,时间和经费约束是确定模型详细程度的主要因素;
- 如果要研究的因素过多,则使用一个"粗"仿真模型或者一个解析模型来识别哪些因素对于系统性能有重大影响。

5.3 Verification of Simulation Computer Programs

OF SIMULATION COMPUTER PROGRAMS In this section we discuss eight techniques that can be used to debug the computer program of a simulation

5.4 Techniques for Increasing Model Validity and Credibility

etailed simulation models5.4.1 Collect High-Quality Information and Data on the System In developing a simulation model <u>show annotation</u>

Collect High-Quality Information and Data on the System:

- Conversations with subject-matter experts
- Observations of the system
- Existing theory
- Relevant results from similar simulation studies
- Experience and intuition of the modelers

Interact with the Manager on a Regular Basis

6. Selecting Input Probability Distributions

6.1 Introductions

hrough 6.7, 6.116.1 INTRODUCTION To carry out a simulation using random inputs such as interarrival times or demand sizes, we have to specify their probability distributions. For example, in the simula-tion <u>show annotation</u>

ade with the simulation results. <mark>If it is possible to collect data on an input</mark> random variable of interest, these data can be used in one of the following approaches to specify a distributionn (in increas-ing order of desir <u>show annotation</u> 如果可能收集到相关的输入随机变量数据,这些数据可用于以下的方法之一,以确 定一个分布:

- 在仿真钟直接使用数据值本身。例如,如果用数据代表服务时间,那么仿真中 一旦需要服务时间就使用一个数据值。这有时叫做 跟踪驱动仿真 (Tracedriven simulation)
- 2. 以某种方式使用数据值本身来定义一个经验分布函数 (seeing <u>Sec 6.2</u>)。如果 这些数据代表服务时间,当仿真中需要服务时间时,就可以从该分布中取样。
- ④ 使用统计推理的标准方法对数据"拟合"一个理论分布形式,例如指数分布或 者泊松分布,并进行假设检验以确定你和的优良度。

ime is needed in the simulation. Two drawbacks of approach 1 are that the simulation can only reproduce what has happened historically and that there is seldom enough data to make all the de-sired simulation runs. Approach 2 avoids these shortco show annotation

6.2 Useful Probability Distributions

6.2.1 Parameterization of Continuous Distributions

e probability density function. However, if the parameters are defined correctly, they can be classified, on the basis of their physical or geometric interpretation, as being one of three basic types: location, scale, or shape parameters. A location parameter g specifie show annotation

可以将参数归类为三种基本类型中的一类:

- Location parametere γ :制定了分布取值范围的横坐标 (absciss, *x* 轴)位置 点,通常是分布的范围的中点(如正态分布的均值 μ) 或者下端点;
- Scale parametere β : 决定了分布范围取值的测量比例,或单位,标准差 σ 是标准正态分布的比例参数;
- Shape parameter α: 决定了相关分布一般类型内的基本形状。α 的变化一般
 会改变分布的性质(如斜歪),较之于位移或者比例参数,其变化是更根本性

的。注:有些分布(如 *指数分布* 或 *正态分布*)没有形状参数,有些分布(如 *贝塔分布*)可以有两个形状参数。

6.2.2 Continuous Distributions

.6.2.2 Continuous Distributions Table 6.3 gives information relevant to simulation modeling applications for 13 con-tinuous distributions. Possible applications are given first to indicate some (certainly not all) uses of the distribution [see Hahn and Shapiro (1994) an show annotation

13 个连续分布

- 均匀分布 Uniform
- 指数分布 Exponential
- 伽马分布 Gamma
- 韦布尔分布 Weibull
- 正态分布 Normal
- 对数正态分布 Lognormal
- 贝塔分布 Beta
- 皮尔逊 V 型分布 Pearson type V
- 皮尔逊 VI 型分布 Pearson type VI
- 对数逻辑斯蒂分布 Log-logistic
- 约翰逊 S_B 分布 Johnson S_B
- 三角分布 Triangular

6.2.3 Discrete Distribution

b x6.2.3 Discrete Distributions The descriptions of the six discrete distributions in Table 6.4 follow the same pattern as for <u>show annotation</u>

六个离散分布:

• 伯努利分布 Bernoulli

- 离散均匀分布 Discrete Uniform
- 二项分布 Binomial
- 几何分布 Geometric
- 负二项分布 negative binomial
- 泊松分布 Poisson

6.2.4 Empirical Distribution

3.6.2.4 Empirical Distributions In some situations we might want to use the observed data themselves to specify directly (in some sense) a distribution, called an empirical distribution, from which random values are g show annotation

经验分布:使用观测数据本身直接(在某种意义下)定义一个分布,这称为经验分 布

cifying empirical distributions. For continuous random variables, the type of empirical distribution that can be defined depends on whether we have the actual values of the individual original observations X1, X2, . . . , Xn rather than only the number of Xi's that fall into each of several specified intervals. (The latter case is called grouped data or data in the form of a histogram.) If the original data are availa show annotation

对于连续随机变量,能定义的经验分布的类型依赖于我们是否具有单个原始观测 X_1, X_2, \dots, X_n 的实际值,而不是只有 X_i 落到若干规定的区间的个数 (后者称为 *分组数据*,或者称为 *直方图形式数据*)

j's that are less than X(i); thi <mark>s is also the way we would like a con-tinuous distribution</mark> function to behave. (See Prob. <u>show annotation</u>

he observations are most dense. The random values generated from this distribution, however, will still be bounded both below (by a0) and above (by ak); see Sec. 8.3.16.In practice, ma <u>show annotation</u>

6.3 Techniques for Assessing Sample Independence

前文讨论的很多统计方法都有一个重要的假设,那就是观测值 X_1, X_2, \dots, X_n 是某个基本分布的独立的(或随机的)样本。例如最大似然估计(Seeing sec 6.5)和 χ^2 检验 (Seeing 6.6.2) 都假设了独立性。

如果关于独立性的假设不成立,那么这些统计方法也许是无效的。

Sec. 6.6.2) assume independence. If the assumption of independence is not satisfied, then these statistical techniques may not be valid . However, even when the data ar <u>show annotation</u>

ively correlated (see Sec. 4.3). We now describe two graphical techniques for informally assessing whether the data X1, X2, . . . , Xn (listed in time order of collection) are independent. The correla-tion plot is a grap show annotation

非正式评估数据 $X_1, X_2, \dots,$ 是否独立 (independent) 的两种图形方法:

- 相关系数图 correlation plot。如果样本相关系数 $\hat{\rho}_j$ 与 0 有很大的数量差异, 那么这是 X_i 非独立的有力证书
- 观测数据 X_1, X_2, \dots 关于 (X_i, X_{i+1}) 点对的散点图 scatter plot。如果 X_i 是 独立的,则期望点 (X_i, X_{i+1}) 会随机地分布在 (X_i, X_{i+1}) 平面的第一象限。如 果*正相关*,则点会呈现一条正斜率的直线;否则呈现 *负斜率* 的直线。

其他的参数(即对 X_i 的分布不做假设)统计检验方法,可以用于正式地检验 X_1, X_2, \dots, X_n 是否独立。

• 冯·诺伊曼比率的排序版本 (A a rank version of Von Neumann's ratio)。一个 潜在的缺陷就是该检验方法假设数据之间没有"粘性"(即对于 $i \neq j, X_i = X_j$)。对于离散数据来说,这个要求一般是不满足的。 queue are positively correlated. There are also several nonparametric (i.e., no assumptions are made about the distributions of the X_i 's) statistical tests that can be used to test formally whether X_1, X_2, \dots, X_n are independent. Bartels (1982) proposes a r show annotation

6.4 Activity 1: Hypothesizing Families of Distributions

SIZING FAMILIES OF DISTRIBUTIONS The first step in selecting a particular input distribution is to decide what general families —e.g., exponential, normal, o <u>show annotation</u>

选择特定输入分布的第一步就是要决定哪些类常用分布形状上与其相似,而不考虑 这些类别的具体参数值。

6.4.1 Summary Statistics

Useful summary statistics

Function	Sample estimate (summary statistic)	Continuous (C) or discrete (D)	Comments
Minimum, maximum	$X_{(1)}, X_{(n)}$	C, D	$[X_{(1)}, X_{(n)}]$ is a rough estimate of the range
Mean μ	$\overline{X}(n)$	C, D	Measure of central tendency
Median x _{0.5}	$\hat{x}_{0.5}(n) = \begin{cases} X_{((n+1)/2)} & \text{if } n \text{ is odd} \\ [X_{(n/2)} + X_{((n/2)+1)}]/2 & \text{if } n \text{ is even} \end{cases}$	C, D	Alternative measure of central tendency
Variance σ^2	$S^2(n)$	C, D	Measure of variability
Coefficient of variation, $cv = \frac{\sqrt{\sigma^2}}{\mu}$	$\widehat{cv}(n) = rac{\sqrt{S^2(n)}}{\overline{X}(n)}$	С	Alternative measure of variability
Lexis ratio, $ au = \frac{\sigma^2}{\mu}$	$\hat{\tau}(n) = rac{S^2(n)}{\overline{X}(n)}$	D	Alternative measure of variability
Skewness, $\nu = \frac{E[(X - \mu)^3]}{(\sigma^2)^{3/2}}$	$\hat{\nu}(n) = \frac{n^2}{(n-1)(n-2)} \frac{\sum_{i=1}^n [X_i - \overline{X}(n)]^3/n}{[S^2(n)]^{3/2}}$	C, D	Measure of symmetry

a continuous distribution. In particular, cv = 1 for the exponential distribution, regardless of the scale parameter β . Thus, cv^{2} (n) being close to

show annotation

s2), since the mean m is zero.] For a discrete distribution, the lexis ratio τ plays the same role that the coefficient of variation does for a continuous distribution. We have found the lexis ratio t show annotation

case of the negative binomial.) The skewness ν is a measure of the symmetry of a distribution. For symmetric distributions like <u>show annotation</u>

6.4.2 Histograms

Quantile Summaries and Box Plots The quantile summary [see, for example, Tukey (1970)] is a synopsis of the sample that is useful for determining whether the underlying probability density function or probability mass function is symmetric or skewed to the right or to the left . It is applicable to either show annotation

6.5 Activity II: Estimation of Parameters

Estimating the parameters from the data

n hypothesized in Activity I, we must somehow specify the values of their parameters in order to have completely specifi ed distributions for possible use in the simulation n. Our IID data X1, X2, ..., show annotation

he quality of an estimator. We shall consider explicitly only one type, maximum-likelihood estimators (MLEs), for three reasons: (1) MLEs have several desirable <u>show annotation</u>

We shall consider explicitly only one type, maximum-likelihood estimators (MLEs), for three reasons:

- MLEs have several desirable properties often not enjoyed by alternative methods of estimation, e.g., least-squares estimators, unbiased estimators, and the method of moments;
- The use of MLEs turns out to be important in justifying the chi-square χ^2 goodness-of-fit test (see Sec. 6.6.2);
- The central idea of maximum-likelihood estimation has a strong intuitive appeal.

distributions.We have said that <mark>MLEs have several desirable statistical properties, some of which are as follows</mark> [see Breiman (1973, pp. 85–88) <u>show annotation</u>

MLE 具有一些很理想的统计性质,包括:

- 对于大多数常用的分布, MLE 是唯一的;也就是说,对于其他任何的 θ 值, L(θ̂) 都严格大于 L(θ);
- 虽然 MLE 不必是无偏的,一般情况下, $\hat{ heta}$ 的渐进分布 $(n o +\infty)$ 的均值等于 heta
- MLE 具有不变性(invariant)。如果对于某个函数 $h, \varphi = h(\theta)$, 那么 φ 的 MLE 为 $h(\hat{\theta})$ (无偏性不是不变性);
- MLE 是渐进正态分布的;
- MLE 是强一致性的,即 $\lim_{n \to +\infty} \hat{ heta} = heta, w. p. 1$

ur purposes. On the other hand, <mark>if the simulation appeared to be sensitive to u, we might seek a better estimate of u;</mark> this would usually entail <u>show annotation</u>

模型不确定性和参数不确定性

nd their associated parameters. When we choose the distributions to use for a simulation model, we generally don't know with absolute certainty whether these are the correct distributions to use, and this lack of complete knowledge results in what we might call model uncertainty . Also, given that cer-tain inpu show annotation ht call model uncertainty. Also, given that cer-tain input distributions have been selected, we typically do not know with com-plete certainty what parameters to use for these distributions, and we might call this parameter uncertainty. (Parameters are typically e <u>show annotation</u>

input probability distributions There have been a number of methods suggested for addressing the problem of input-model uncertainty, including the following [see Barton (2012) and Henderson (2003)] :• Bayesian model averaging [Ch <u>show annotation</u>

There have been a number of methods suggested for addressing the problem of input-model uncertainty, including the following [see Barton (2012) and Henderson (2003)]:

- 贝叶斯模型平均值方法 (Bayesian model averaging) [Chick (2001), Zouaoui and Wilson (2003, 2004)]
- Delta 方法 (Delta-method approaches) [Cheng and Holland (1997, 1998, 2004)]
- 元模型辅助的自举方法 (Metamodel-assisted bootstrapping) [Barton et al. (2013), Chapter 12]; see Cheng (2006) and Efron and Tibshirani (1993) for a discussion of bootstrap resampling
- 非参数的自举方法 (Nonparametic bootstrapping) [Barton and Schruben (1993, 2001)]
- 基于随机效用模型的快速方法 (Quick method based on a random-effects model) [Ankenman and Nelson (2012)]

del [Ankenman and Nelson (2012)] <mark>Unfortunately, most of these methods are reasonably complicated and make assumptions that may not always be satisfi ed in practice</mark> [see Barton (2012) and Bar <u>show annotation</u>

6.6 Activity III: Determining How Representative The Fitted Distributions Are

IVE THE FITTED DISTRIBUTIONS ARE After determining one or more probability distributions that might fit our observed data in Activities I and II, we must now closely examine these distributions to see how well they represent the true underlying distribution for our data . If several of Law01323_ch06_32 show annotation

评估拟合的分布的质量的方法:

- 启发式方法(Heuristic procedures)
- 拟合优良度假设检验法(Goodness-of-fit hypothesis tests)

五个启发式方法:

- 密度直方图和频率比较
- 分布函数差图
- 概率图(quantile-quantile(Q-Q)plot、probability-probability(P-P)plot)

nterval for j 5 1, 2, . . . , k. For discrete data, a frequency comparison is a graphical comparison of a histo-gram of the data with the mass function p[^] (x) of a fi tted distribution . Let hj be the observed proport <u>show annotation</u>

a set (see Prob. 6.15), a <mark>quantile–quantile (Q–Q) plot</mark> (see Sec. 6.4.3 for the d

show annotation

all to mod-erate sample sizes. A probability—probability (P—P) plot is a graph of the model probabi show annotation

6.6.2 Goodness-of-Fit Tests

ons.6.6.2 Goodness-of-Fit Tests A goodness-of-fit test is a statistical hypothesis test (see Sec. 4.5) that is used to assess formally whether the observations X_1, X_2, \dots, X_n are an independent sample from a particular

distribution with distribution function \hat{F} . That is, a goodness-of-fit tes show annotation

拟合优良度检验可以用于检验下面的原假设:

 $H_0: X_i$ 是独立同分布的随机变量,具有分布函数 \hat{F}

Note: 未能拒绝 H_0 并不意味着"接受 H_0 为真"

properties of these tests. First, failure to reject H0 should not be interpreted as "accepting H0 as being true." Law01323_ch06_324-392.indd Page <u>show annotation</u>

具体的拟合优良度检验方法:

- χ² 检验
- 科尔莫戈罗夫-斯米尔洛夫检验
- 安德森-达林检验
- 泊松过程检验

to specify the intervals. Kolmogorov-Smirnov (K-S) tests for goodness of fit, on the other hand, compare an empirical distribution function with the distribution function F[^] of the hypothesized distribution. As we shall see, K-S tests do

show annotation

正如我们刚刚看到的一样, χ^2 检验可以看做将数据直方图与拟合分布的密度函数或 者质量函数进行比较正式的对比。在连续情况下应用 χ^2 检验的一个实际困难,即如 何确定区间。另一方面,拟合优良度的科尔莫戈罗夫-斯米尔洛夫(Kolmogorov-Smirnor, K-S)检验将经验分布函数与假设分布的分布函数 \hat{F} 作对比。正如我们将 要看到的一样,K-S 检验不需要以任何形式对数据分组,因此没有任何信息丢失; 这些解决了区间划分这个麻烦问题。K-S检验另一个优点就是,它们对于任意样本大 小 n(在所有参数都已知的情况下)都是(严格)有效的,而x检验只在渐近意义上 有效。最后,在很多可选分布的情况下,K-S检验能力比 χ^2 检验强;参见,例如, Stephens(1974)。 input probability distributions Nevertheless, K-S tests do have some drawbacks, at least at present. Most seri-ously, their range of applicability is more limited than that for chi-square tests. First, for discrete data, the re <u>show annotation</u>

6.10 Specifying Multivariate Distributions Correlations, and Stochastic Processes

ATIONS, AND STOCHASTIC PROCESSES So far in this chapter we have considered only the specifi cation and estimation of the distribution of a single, univariate random variable at a time. If the simulation model needs show annotation

本章目前为止,讨论的知识每次确定和估计单个、单随机变量的分布。如果仿真模型需要的输入只是标量随机变量,而且如果他们在模型之间相互独立,那么对每个 输入重复地使用至今讨论的方法就足够了。

ted by most simulation packages. There are systems, however, in which the input random variables are statisti-cally related to each other in some way: • Some of the input random show annotation

然而,有些系统的输入随机变量之间以某种方式统计相关:

有些输入随机变量共同形成一个随机向量,具有某种多元(或联合)概率分
 布,需要建模着确定。

Some of the input random variables together form a random vector with some multivariate (or joint) probability distribution (see <u>Sec. 4.2</u>) to be specified by the modeler.

6.10.2 Specifying Arbitrary Marginal Distributions and Correlations

compose an input random vector. In each of these cases, the fi tted member of the multivariate distribution family involved (normal, lognormal, Johnson, or Bézier) determined the correlation between pairs of the component random variables in the vector, as well as their marginal distributions; it also imposed a more general <u>show annotation</u>

greater fl exibility than that. We may want to allow for pos-sible correlation between various pairs of input random variables to our simulation model, yet not impose an overall multi show annotation

6.10.3 Specifying Stochastic Process

ProcessesAs mentioned earlier, there are situations in which a sequence of input random vari-ables on the same phenomenon are appropriately modeled as being draws from the same (marginal) distribution, yet might exhibit some autocorrelation between themselves within the sequence . For instance, if {X1, X2, show annotation

平稳随机过程:

- AR 和 ARMA 过程
- 伽马过程
- ARTA 过程
- VARTA 过程

om vectors.AR and ARMA Processes <mark>Standard autoregressive (AR) or</mark> autoregressive moving-average (ARMA) models, developed in Box et al. (<u>show annotation</u>

6.11 Selecting a Distribution in the Absence of Data

TRIBUTION IN THE ABSENCE OF DATA In some simulation studies it may not be possible to collect data on the random variables of interest, so the techniques of Secs. 6.4 through 6.6 are not applicable to the problem of selecting corresponding probability distributions. For example, if the system bei show annotation

6.12 Models of Arrival Processes

at or before time t for t \$ 0. We call the stochastic process {N(t), t \$ 0} an arrival process since, for our purposes, the events <u>show annotation</u>

Arrival process

令 $N(t) = \max\{i : t_i \leq t\}$ 为t时刻或以前发生的事件的个数。称随机过程 { $N(t), t \geq 0$ }为到达过程,其中 $A_i = t_i - t_{i-1}$ ($i = 1, 2, \cdots$)是第i - 1个 和第i个顾客之间的到达间隔时间。

6.12.1 Poisson Processes

son ProcessesIn this section we <mark>defi ne a Poisson process, state some of its important properties</mark> , and in the course of doing so <u>show annotation</u>

Definition:

- 顾客每次到达一个;
- N(t+s) N(t) (在时间区间 (t,T+s] 上到达的个数) 独立于 {N(u), 0 ≤ u ≤ t};
- 对于所有 $t, s \ge 0$, 分布 N(t + s) N(t) 独立于 t。

aw01323_pagefileschapter six 381 the number of arrivals in the interval (t, t 1 s] is independent of the number of arrivals in the earlier time interval [0, t] and also of the times at which these arrivals occur . This property could be violate

show annotation

6.2 below and then Example 6.4.) The following theorem, proved in Çinlar (1975, pp. 74–7

show annotation

定理1: 如果 { $N(t), t \ge 0$ } 是一个泊松过程,那么在长度为 *s* 的任意时间区间上到达的数目是参数为 λs (λ 为正实数) 的柏松随机变量。即:

$$P\left[N(t+s)-N(t)=k
ight]=rac{e^{-\lambda s}(\lambda s)^k}{k!} \hspace{1em} k=0,1,\cdots, ext{and} \hspace{1em} t,s\geq 0$$

因此, $\mathbb{E}[N(s)] = \lambda s$ (See Sec 6.2.3), 且, $\mathbb{E}[N(1)] = \lambda$ 。因此 λ 是长度为 1 的任意区间内到达的期望数目,称之为过程的速率。

定理2: 如果 { $N(t), t \ge 0$ } 是速率为 λ 的泊松过程,那么它对应的到达间隔时间 A_1, A_2, \dots 为 IID 指数随机变量,均值为 $1/\lambda$.

6.12.2 Nonstationary Poisson Processes

本节讨论普遍使用的具有时变到达速率的到达过程的模型。

随机过程 { $N(t), t \ge 0$ } 是一个非平稳泊松过程,如果:

- 顾客每次到达一个;
- N(t+s) N(t) 独立于 { $N(u), 0 \le u \le 0$ }

令 $\Lambda(t) = \mathbb{E}[N(t)]$,对于所有 $t \ge 0$,如果对于特定的 t 值 $\Lambda(t)$ 可微 (differentiable),将 $\lambda(t)$ 为:

$$\lambda(t) = rac{\mathrm{d}}{\mathrm{d}\,t}\Lambda(t)\,.$$

直观来看,在到达的期望数目大的区间 $\lambda(t)$ 将会很大,分别称 $\Lambda(t)$ 和 $\lambda(t)$ 为非平 稳泊松过程的 期望函数 和 速率函数 。

d number of arrivals is large. We call L(t) and I(t) the expectation function and the rate function, re-spectively, for the nonstationary Poisson process. The following theorem shows <u>show annotation</u>

定理3:如果 { $N(t), t \ge 0$ } 为具有连续的期望函数 $\Lambda(t)$ 的非平稳泊松过程,那么:

$$P[N(t+s)-N(t)=k] = rac{e^{-b(t,s)}[b(t,s)]^k}{k!} \quad ext{for } k=0,1,2,\dots ext{ and } t,s \geq 0$$

其中, $b(t,s) = \Lambda(t+s) - \Lambda(t) = \int_{t}^{t+s} \lambda(y) dy$, 对于 [t,t+s] 上所有但有限多个 点, 如果 $d(\Lambda(t))/dt$ 在 [t,t+s] 上是有界的且如果 $d(\Lambda(t))/dt$ 存在并连续, 那么 最后一个等式成立。

rameter depends on both t and s. T H E O R E M 6 . 3 . If {N(t), t \$ 0} is a nonstationary Poisson process with continuous expectation function L(t), then P[N(t 1 s) 2 N(t) 5 k] 5 e2 b(t show annotation

ite simple and fairly fl exible, <mark>it does require somewhat arbitrary judgment</mark> about the boundaries and widths of the constant-rate time intervals . Other methods ^ (t)0.5011:00 <u>show annotation</u>

6.12.3 Batch Arrivals

process.6.12.3 Batch Arrivals For some real-world systems, customers arrive in batches, or groups, so that property 1 of the Poisson process and of the nonstationary Poisson process is violated . For example, people arrivin <u>show annotation</u>

t model such an arrival process. Let N(t) now be the number of batches of individual customers that have arrived by time t. By applying the techniques dis show annotation

can be made more precise. If X(t) is the total number of individual customers to arrive by time t, and if B_i is the number of customers in the *i*th batch, then X(t) is given by X(t) 5 show annotation

6.13 Assessing the Homogeneity of Different Data Sets

有时分析人员对一个随机现象独立地收集了 k 组观测值, 他想知道这些数据集是否 同质从而能合并。

MOGENEITY OF DIFFERENT DATA SETS Sometimes an analyst collects k sets of observations on a random phenomenon in-dependently and would like to know whether these data sets are homogeneous and thus can be merged . For example, it might be of in

show annotation

有时分析人员对一个随机现象独立地收集了 k 组观测值,他想知道这些数据集是 否同质从而能合并。

本节讨论同质性的克鲁斯卡-沃尔斯 (Kruskal-Wallis) 假设检验。由于对数据的分布 不做假设,这是非参数化检验。

ion is needed. In this section, <mark>we discuss the Kruskal-Wallis hypothe-sis test for homogeneity.</mark> It is a nonparametric test sinc show annotation

7. Random-Number Generators

7.1 Introduction

3.2, 7.4.1, 7.4.37.1INTRODUCTION A simulation of any system or process in which there are inherently random components requires a method of generating or obtaining numbers that are random, in some sense. For

example, the queueing a <u>show annotation</u>

in executing simulation models. So as to avoid speaking of "generating random variables," which would not be strictly correct since a random variable is defined in mathematical probability theory as a function satisfying certain conditions, we will adopt more precise terminology and speak of "generating random variates." This entire chapter is devoted t <u>show annotation</u> 本章致力于讨论由 [0,1] 区间上的均匀分布产生随机变数的方法。

of "generating random variates." This entire chapter is devoted to methods of generating random variates from the uniform distribution on the interval [0, 1]; this distribution was denoted by <u>show annotation</u>

uch independent random numbers. This prominent role of the U(0, 1) distribution stems from the fact that random vari-ates from all other distributions (normal, gamma, binomial, etc.) and realizations of various random processes (e.g., a nonstationary Poisson process) can be obtained by transforming IID random numbers in a way determined by the desired distribution or process. This chapter discusses ways to show annotation

所有其他分布(正态分布、伽马分布、二项分布等)的随机变数以及各种随机过 程(如非平稳泊松过程)的实现都可以通过按所要求的分布或过程确定的方法对 IID 随机数进行变换的方法来得到

with the way com-puters work. One possibility would be to hook up an electronic random-number machine, such as ERNIE, directly to the computer. This has several disadvantages, chiefl y that we could not reproduce a previously generated random-number stream exactly. (The desirability of being able show annotation

01323_pagefileschapter seven 395 Intuitively the midsquare method seems to provide a good scrambling of one number to obtain the next, and so we might think that such a haphazard rule would provide a fairly good way of generating random numbers . In fact, it does not work very <u>show annotation</u>

y of generating random numbers. In fact, it does not work very well at all. One serious problem (among others) is that it has a strong tendency to degenerate fairly rapidly to zero, where it will stay forever. (Continue Table 7.1 for just show annotation

缺陷;有很强的非常快速地退化到0点趋向。

o one number to obtain the next. <mark>A more fundamental objection to the midsquare method is that it is not "random" at all, in the sense of being unpredictable.</mark> Indeed, if we know one number, <u>show annotation</u>

ature of truly random numbers. (Sometimes arithmetic generators are called pseudorandom, an awkward term that we avoid, even though it is probably more accurate.) Indeed, in an oft-quoted quip, <u>show annotation</u>

s shared by Lehmer (1951), who developed what is probably still the most widely used class of techniques for random-number TABLE 7.1The midsquare method i <u>show annotation</u>

ers that arithmetic generators, if designed carefully, can produce numbers that appear to be independent draws from the U(0, 1) distribution, in that they pass a series of statistical tests (see Sec. 7.4). This is a usefu <u>show annotation</u>

numbers," to which we subscribe. <mark>A "good" arithmetic random-number</mark> generator should possess several properties: 1. Above all, the numbers produ <u>show annotation</u>

一个"好的"算数随机数发生器应当具备以下属性:

- 所产生的数字应当呈现出在 [0,1] 区间上是均匀分布的,并且相互之间不应存
 在相关性;否则仿真的结果将会完全无效;
- 从实用的角度来说,我们希望发生器具有速度快,并且不需要过的的存储空间;
- 我们要能严格地重复生成一个给定的随机数流,原因至少有二。1)这有时可 能使得调试或校验计算机程序更容易;2)我们可能需要使用一组同样的随机

数仿真不同的系统以便得到更为精确的比较;

- 发生器应当具备有易于长生分开的随机数流;
- 我们希望发生器是方便的,即对所有标准的编译器和计算机来说,产生相同的随机数序列。

sed in other parts of Chap. 11. The ability to create separate streams for a generator is facilitated if there is an effi cient way to jump from the ith random number to the (i 1 k)th random number for large values of k. 5. We would like the generator show annotation

are almost universally met. Furthermore, most generators now have the facility for multiple streams in some way, especially those generators included in modern simulation packages, satisfying point 4 . Unfortunately, there are also show annotation

n packages, satisfying point 4. Unfortunately, there are also some generators that fail to satisfy the uniformity and independence criteria of point 1 above, which are absolutely necessary if one hopes to obtain correct simulation results. For example, L'Ecuyer and Simar <u>show annotation</u>

许多发生器为满足均匀性和独立性准则,如果人们期望得到正确的仿真结果,则 这是必要的。

7.2 Linear Congruential Generators

.2LINEAR CONGRUENTIAL GENERATORS Many random-number generators in use today are linear congruential generators (LCGs), introduced by Lehmer (1951). A sequence of integers Z1, Z2, show annotation

线性同余发生器

整数 Z_1, Z_2, \cdots 序列由下面的迭代公式定义:

 $Z_i = (aZ_{i-1} + c)$ (对 m 取模)

其中, m 表示 模数 (modulus), a 表示 乘子 multiplier, c 表示增量 increment, 以 及 Z_0 表示种子或初始值 seed 都是非负整数。并且, $0 < m, a < m, c < m, Z_0 < m$

0 , m, a , m, c , m, and Z0 , m. Immediately, two objections could be raised against LCGs. The fi rst objection is one co <u>show annotation</u>

线性同余发生器 LCG 的两个缺点:

- 所有(伪)随机数发生器共同的问题,即上述公式定义的 *Z_i* 根本不是真正的 随机数。
- U_i 的取值只能是有理值 $0, 1/m, 2/m, \dots, (m-1)/m$, 事实上, U_i 实际只会 取其中的一部分,这依赖于 m, a, c, Z_0 之间的规定,以及 m 浮点除的性质。

cycle repeats itself endlessly. The length of a cycle is called the period of a generator. For LCGs, Zi depends only on the previous integer Z_{i_1} , and since $0 \le Z_i \le m - 1$, it is clear that the period is at most m; if it is in fact m, the LCG is said to have full period. (The LCG in Example 7.2 has ful show annotation

满周期

#full-period

of the seed for this generator. Since large-scale simulation projects can use millions of random numbers, it is manifestly desirable to have LCGs with long periods. Furthermore, it is comforting t show annotation

ng to have full-period LCGs, <mark>since we are assured that every integer between</mark> 0 and m 2 1 will occur exactly once in each cycle, which should contribute to

the unifor-mity of the Ui's. (Even full-period LCGs, however <u>show annotation</u>

因为大规榄系统仿真项目可能使用几百万个随机数,显然希望具有长周期的 LCG, 进一步,最好是 *满周期*,因为我们就能保证在每个循环中 $0 \sim m-1$ 之间的每个整 数將份好出现一次,这会有助于 U_i 的均匀性(但是,即使是满周期的 LCG,按一 个循环的段来说,也可能呈现出非均匀性。

例如,假如我们只生成 m/2 个连续的 Z_i , 它们在可能的取值 $0, 1, \dots, m-1$ 序列中有可能留下大间隙。

所以,了解如何选择参数 *m*、*a*和 *c*。使得相应的 LCG 具有满周期是有用的。下面 的定理给出了这种特征描述。

定理7.1 由 LCG 的定义的 LCG 具有全周期,当且仅当下列三个条件成立:

1. m 和 *c* 能够同时被整除的正整数只有 1, 即 *m* 与 *c* 互质;

2. 如果 q 为整除 m 的素数(只能被 1 和其自身整除),则 q 能整除 a-1;

3. 如果 4 整除 m, 则 4 整除 a-1。

定理7.1中的条件(a)通常表述为"m与c互质。"

日者满周期(盐老至心具长周期)示具好的iCC 的一个所杀胡的从

中指出的,希望好的统斗性料

LCG will have full period. T he following theorem, proved by Hull and Dobell (1962), gives such a characterization .T H E O R E M 7 . 1 . The show annotation

LCG; as indicated in Sec. 7.1, we also want good statistical properties (such as apparent independence), computational and storage effi ciency, reproducibility, facilities for separate streams, and portability (see Sec. 7.2.2). Reproducibility is simple, f <u>show annotation</u> 在上述定理中的条件 1,我们知道 c > 0 (称为混合 LCG)与 c = 0 (称为乘 LCG)将会有不同的表现。

7.2.1 Mixed Generators

iles400 random-number generators most computers and compilers have 32bit words, the leftmost bit being a sign bit, so b 5 31 and m 5 231 . 2.1 billion . Furthermore, choosing m 5 2b d <u>show annotation</u>

231 . 2.1 billion. Furthermore, <mark>choosing m 5 2b does allow us to avoid explicit division by m on most computers by taking advantage of integer overfl ow . The larg-est integer that can <u>show annotation</u></mark>

7.2.2 Multiplicative generators

#乘法发生器 #乘

#乘LCG

7.2.2 Multiplicative Generators Multiplicative LCGs are advantageous in that the addition of c is not needed, but they cannot have full period since condition (a) of Theorem 7.1 cannot be satisfied d (because, for example, m is po show annotation

m and c 5 0). As we shall see, however, it is possible to obtain period m - 1 if m and a are chosen carefully .As with mixed generators, it's show annotation

if m and a are chosen carefully. As with mixed generators, it's still computationally efficient to choose $m = 2^b$ and thus avoid explicit division . However, it can be shown [see, show annotation]

与混合发生器一样,取 $m = 2^b$, 计算效率仍然高且避免显式除法。但是在这种情况下,其周期最多为 2^{b-2} , 也就是说, Z_i 所能取得数值的个数仅仅为 $0 \sim m - 1$ 之间

所有整数个数的四分之一。

事实上, 如果 Z_0 为奇数, 对于某个 $k = 0, 1, \dots, a$ 可以表示成 8k + 3 或者 8k + 5 这样的形式, 则周期为 2^{b-2} 。

tained as values for the Zi's. (In fact, the period is 2^{b-2} if Z_0 is odd and a is of the form 8k + 3 or 8k + 5 for some $k = 0, 1, \cdots$) Furthermore, we generally s show annotation

properties can be induced. T he generator usually known as RANDU is of this form ($m = 2^{31}, a = 2^{16} + 3 = 65539, c = 0$) and has been shown to have very undesirable statistical properties (see Sec. 7.4). Even if one does not choos show annotation

普遍熟知的发生器 RANDU 就是这样的形式,即 $m = 2^{31}, a = 2^{16} + 3 = 65539, c = 0$,而且已经证明,这个发生器具有 不尽如人意的统计特性。

result-ing possibility of gaps. Because of these difficulties associated with choosing $m = 2^b$ in multiplicative LCGs, attention was paid to finding other ways of specifying m. Such a method, which has pr show annotation

由于在乘 LCG 中选 $m = 2^b$ 带来的这些困难,人们将注意力放在寻找其他方法来规 定 m 值。不是令 $m = 2^b$, 而建议 m 取小于 2^b 的最大素数。例如,在 b = 31 的情 形,小于 2^{31} 的最大素数非常容易得到 2^{31} —1 = 2147483647。现在,对于素数m, 可以证明,如果 a 为模 m 的 <u>素元</u> (primitive element modulo),则周期为 m—1, 即使 $a^l - 1$ 能被 m 整除的最小整数是 l = m - 1。

对于以这种方法选择的 m 和 a,我们得到每个整数 1, 2,…, m-1 在每个循环内 正好出现一次,所以 Z_0 可以是 $1 \sim m-1$ 的任意整数,得到的周期仍是 m-1。这 称为*素数取模乘 LCG* (prime modulus multiplicative LCGs, PMMLCGs)。

iod of m 2 1 will still result. These are called prime modulus multiplicative LCGs (PMMLCGs) .Two issues immediately arise co <u>show annotation</u>

s multiplicative LCGs (PMMLCGs). <mark>Two issues immediately arise concerning PMMLCGs:</mark> (1) How does one obtain a prim <u>show annotation</u>

- 与 PMMLCGs 相关的两个问题随即产生:
 - 1. 如何得到模 *m* 的素元? 将通过讨论下面两个广泛使用的 PMMLCG 在本质上解 决这一点;
 - 2. 由于不选 $m = 2^{b}$,再也不能直接使用 <u>整数溢出</u> 机制来达到取 m 模除法的效果。这种情况下,避免显式除法的方法也适用一类溢出,称为 *仿真除* (simulated division)。

below, uses simulated division. Considerable work has been directed toward identifying good multipliers a for PMMLCGs t hat are primitive elements m show annotation

对于 PMMLCG 来说,重要的是取模 $m^* = 2^{31} - 1$ 的素元的好乘子 a,从而得到 周期为 $m^* - 1$ 。

er of fairly stringent criteria. ==Two particular values of a that have been widely used for the modulus m* are $a_1 = 7^5 = 16807$ and $a_2 = 630, 360, 016$, both of which are primitive elements modulo m== . [However, neither value of a w*

show annotation

对于模数 m^* ,已经取得广泛使用的 a 的两个特定乘数值为 $a_1 = 7^5 = 16807$ 和 $a_2 = 630, 360, 016$,这两个都是取模 m^* 的素元。

然而,并没有发现哪个 a 值是其中最好的, see Sec 7.4.2.

he best (see Sec. 7.4.2).] The multiplier a1 was originally suggested by Lewis, Goodman, and Miller (1969), and it was used by Schrage (1979) in a clever FORTRAN implementation using simulated division. The importance of Schrage's code was that it provided at that time a reasonably good and portable random-number generator. The multiplier a2, suggested ori 乘子 a_1 采用仿真除法在一个聪明的 FORTRAN 实现中使用了 a_1 ,其代码重要性在 于,他在当时提供恶意个相当好的且便于移植的随机数发生器。

numbers is not too large. However, many experts [see, e.g., L'Ecuyer, Simard, Chen, and Kelton (2002) and Gentle (2010, p. 21)] recommend that LCGs with a modulus of around 231 should no longer be used as the randomnumber generator in a general-purpose software package (e.g., for discreteevent s

<u>show annotation</u>

discrete-event simula-tion). Not only can the period of the generator be exhausted in a few minutes on many computers, but, more importantly, the relatively poor statistical properties of these generators can bias simulation results for sample sizes that are much smaller than the period of the generator . For example, L'Ecuyer and Sima show annotation

不建议作为通用仿真软件包中的随机数发生器:

- 1. 一个周期的随机数在很多计算机上往往在几分钟内就被用尽;
- 这些随机数相对较差的统计性会使得仿真结果偏离样本长度,后者较发生器周 期要小很多。

7.3 Other Kinds of Generators

red.7.30THER KINDS OF GENERATORS Although LCGs are probably the most widely used and best understood kind of random-number generator, there are many alternative types. (We have already seen one alte <u>show annotation</u>

开发其他类型发生器的主要目的是要得到更长的周期和更好的统计特性。

7.3.1 More General Congruences

LCGs can be thought of as a special case of generators defined by:

 $Z_i = g(Z_{i-1}, Z_{i-2}, \cdots) (ext{mod}\ m)$

where g is a fixed deterministic function of previous Z_j 's.

两种推广:

- 二次同余发生器 (quadratic congruential generator)
- 多重递归发生器 (multiple recursive generator, MRG)

a more detailed discussion. One obvious generalization of LCGs would be to let g(Zi21, Zi22, . . .) 5 a9Z2i21 1 aZi21 1 c, which produces a quadratic congruential generator . A special case that has receiv <u>show annotation</u>

ators is at most m, as for LCGs. <mark>A different choice of the function g is to maintain linearity but to use earlier Zj's; this gives rise to generators called multiple recursive generators (MRGs) and defi ned by g(Zi21, Zi 2 2, <u>show annotation</u></mark>

7.3.2 Composite Generators

7.12.7.3.2 Composite Generators Several researchers have developed methods that take two or more separate genera-tors and combine them in some way to generate the fi nal random numbers. It is hoped that this composi

show annotation

复合发生器:将两个或多个单独的发生器以某种方式结合起来生成最终的随机数。 期望这种组合发生器较之组成它的任何单个发生器有更长的周期,更好地统计性 能。*缺点*:代价大于单独使用一个发生器。

simple generators composing it. The disadvantage in using a composite generator is, of course, that the cost of obtaining each U_i is more than that of using one of the simple generators alone. Perhaps the earliest kind of com show annotation

andom numbers (see Sec. 11.2.3). Wichmann and Hill (1982) proposed the following idea for combining three generators, again striving for long period, portability, speed, and usability on small computers (as well as statistical adequac show annotation

7,3,3 Feedback Shift Register Generators

• 线性反馈移位寄存器 (Linear feedback shift register, LFSR)

by dividing the Wi's by 16 5 24. The original motivation for suggesting that the bi's be used as a source of U(0, 1) random numbers came from the observation that the recurrence given by Eq. (7.4) can be implemented on a binary computer using a switching circuit called a linear feedback shift register (LFSR). This is an array of q bits tha show annotation

. . , 15 are given in Table 7.3. Unfortunately, LFSR generators are known to have statistical defi ciencies, as discussed by Matsumoto and K show annotation

性能差

- 广义反馈移位寄存器 (Generalized feedback shift register, GFSR)
- 旋转的广义反馈移位寄存器 (twisted generalized feedback shift register, TGFSR)

the same as recurrence (7.7).] With suitable choices for r, q, and A, a TGFSR generator can have a maximum period of $2^{ql} - 1$ as compared with $2^q - 1$ for a GFSR generator (both require ql bits to store the state of the generator). Matsumoto and Kurita (1994) di show annotation

7.4 Testing Random Number Generators

TESTING RANDOM-NUMBER GENERATORS As we have seen in Secs. 7.1 through 7.3, all random-number generators currently used in computer simulation are actually completely deterministic. Thus, we can only hope that the U_i 's generated appear as if they were IID U(0,1) random variates . In this section we discuss sev show annotation

art of the avail-able software. Before such a generator is actually used in a simulation, we strongly recommend that one identify exactly what kind of generator it is and what its nu-merical parameters are. Unless a generator is one of th show annotation

7.4.1Empirical Tests

's at all.7.4.1 Empirical Tests Perhaps the most direct way to test a generator is to use it to generate some U'_i s, which are then examined statistically to see how closely they resemble IID U(0,1) random variates . We discuss four such empir

show annotation

本节将讨论四种实验检验方法,统计检查随机数类似 U(0,1) 随机数的紧密程度。

- 1. 所有参数已知的 χ^2 检验的特殊情形;
- 2. 连续检验,实际上是把 χ^2 检验推广到高维;
- 游程(或上升游程)检验,是对独立性假设的更直接检验(实际上,它只检验 独立性,即把特别是不检验均匀性);
- 4. 用直接的方法以评价所产生的 U_i 是否存在可辨别的相关性: 对于某个 l,只计算其滞后 $j = 1, 2, \cdots$ 的滞后 j 的相关系数的估计。

uniformity in three dimensions. Why should we care about this kind of uniformity in higher dimensions? If the individual U_i 's are correlated, the distribution of the d-vectors U_i will deviate from d-dimensional uniformity; thus, the serial test provides an indirect check on the assumption that the individual U_i's are independent. For example, if adjacent Ui's te <u>show annotation</u>

ently fi ne division of [0, 1].) The third empirical test we consider, the runs (or runs-up) test, is a more di-rect test of the independence assumption. (In fact, it is a test o show annotation

e autocorrelation at these lags. <mark>As mentioned above, these are just four of</mark> the many possible empirical tests. For example, the Kolmogorov-Smirnov test discussed in Sec. 6.6.2 (for the case with all paramete

show annotation

niformity on the unit cube. <mark>RANDU is a fatally fl awed generator, due primarily to its utter failure in three dimensions;</mark> we shall see why in Sec. 7.4.2. <u>show annotation</u>

we shall see why in Sec. 7.4.2. One potential disadvantage of empirical tests is that they are only local; i.e., only that segment of a cycle (for LCGs, for example) that was actually used to generate the Ui's for the test is examined, so we cannot say anything about how the generator might perform in other segments of the cycle. On the other hand, this I <u>show annotation</u>

<mark>实验检验的缺陷</mark>:他们总是局部的,也就是说,仅仅对循环中的一段进行了检验, 这一段实际上是用于产生检验用的 *U_i*。

7.4.2 Theoretical Tests

s known.7.4.2 Theoretical Tests We now discuss theoretical tests for randomnumber generators. Since these tests are quite sophisticated and mathematically complex, we shall describe them somewhat qualitatively; for detailed accounts see F <u>show annotation</u> Ecuyer (1998, pp. 106–114). A as mentioned earlier, theoretical tests do not re-quire that we generate any U_i 's at all but are a priori in that they indicate how well a generator can perform by looking at its structure and defining constant s. Theo-retical tests also diffe show annotation

ucture and defi ning constants. Theoretical tests also differ from empirical tests in that they are global; i.e., a genera-tor's behavior over its entire cycle is examined. As we mentioned at the end <u>show annotation</u>

or global tests are preferable; global tests have a natural appeal but do not generally indicate how well a specifi c segment of a cycle will behave .It is sometimes possible to com <u>show annotation</u>

7.4.3 Some General Observations on Testing

is absolutely "the best." One piece of advice that is often offered, however, is that a random-number generator should be tested in a way that is consistent with its intended use. This would entail, for example, <u>show annotation</u>

8. Generating Random Variables

8.1 Introduction

ing: 8.1 and 8.28.1 INTRODUCTION <mark>A simulation that has any random aspects at all must involve sampling, or generat-ing, random variates from probability distributions.</mark> As in Chap. 7, we use the phrase <u>show annotation</u>

from the desired distribution. These distri-butions are often specifi ed as a result of fi tting some appropriate distributional form, e.g., exponential, gamma, or Poisson, to observed data, as discussed in Chap. 6 . In this

chapter we assume that show annotation

values of the parameters), and we address the issue of how we can generate random variates with this distribution in order to run the simulation model. For example, the queueing-type <u>show annotation</u>

As we shall see in this chapter, the basic ingredient needed for every method of generating random variates from any distribution or random process is a source of IID U(0,1) random variates . For this reason it is essentia show annotation

从任意分布或随机过程中生成随机变量的每一种方法所需要的基本要素是一个独 立同分布的U(0,1)随机变量源。

rform adequately (see Chap. 7). Without an ac-ceptable random-number generator, it is impossible to generate random variates correctly from any distribution . In the rest of this chapter, w <u>show annotation</u>

from a given distribution, and several factors should be considered when choosing which algorithm to use in a particular simulation study. Unfortunately, these different factors often confl ict with each other, so the analyst's judg-ment of which algorithm to use must involve a number of tradeoffs . All we can do here is raise so <u>show annotation</u>

选择什么算法来生成随机变量,应该考虑多种因素:

- 准确性
- 效率
- 复杂性
- 技术

some of the pertinent questions. The first issue is exactness. We feel that, if possible, one should use an algo-rithm that results in random variates with exactly the desired distribution, within the unavoidable external limitations of machine accuracy and exactness of the U(0, 1) random-number generator. Effi cient and exact algorithm

I prefer to use an exact method. Given that we have a choice, then, of alternative exact algorithms, we would clearly like to use one that is effi cient, in terms of both storage space and execution time. Some algorithms require sto

<u>show annotation</u>

come an important consideration. <mark>A somewhat subjective issue in choosing an algorithm is its overall complexity, including conceptual as well as implementational factors.</mark> One must ask whether the potent <u>show annotation</u>

r a "one-time" simulation model. Finally, there are a few issues of a more technical nature. Some algorithms rely on a source of random variates from distributions other than U(0, 1), which is unde-sirable, other things being equal. Another technical issue is that show annotation

es428 generating random variates used variance-reduction techniques (common random numbers and antithetic vari-ates) require synchronization of the basic U(0, 1) input random variates used in the simulation of the system(s) under study, and this synchronization is more easily accomplished for certain types of random-variate generation algorithms . In particu-lar, the general show annotation

ss two more specialized topics: generat-ing correlated random variates and generating realizations of both stationary and non-stationary arrival processes .8.2 GENERAL APPROACHES TO GENER <u>show annotation</u>

8.2 General Approaches to Generating Random Variates

ES TO GENERATING RANDOM VARIATES There are many techniques for generating random variates, and the particular algorithm used must, of course, depend on the distribution from which we wish to generate; however, nearly all these techniques can be classifi ed according to their theoretical basis . In this section we discuss the show annotation

8.2.1 Inverse Transform

the inverse of the function F. Then an algorithm for generating a random variate X having distribution function F is as follows (recall that , \sim is read "is distributed as"): 1. Generate U , U(0, 1).2. Ret show annotation

假定要产生一个连续随机变量 *X* (See Sec 4.2), *X* 有分布函数 *F*, 当 0 < F(x) < 1 时, *F* 是连续的且严格递增的(这意味着如果 $x_1 < x_2$ 且 $0 < F(x_1) \le F(x_2) < 1$,则事实上 $F(x_1) < F(x_2)$)。

令 F^{-1} 表示函数 F 的逆,则生成具有分布函数 F 的随机变量 X 的算法如下:

- 1. 产生 $U \sim U(0,1)$
- 2. 返回 $X = F^{-1}(U)$
- 通用反变换法 (General inverse-transform method)

eads to the negative variate X2. To show that the value X returned by the above algorithm, called the general inverse-transform method, has the desired distribution F, show annotation

• Transformer uniform distribution to exponential distribution

EXAMPLE 8.1. Let X have the exponential distribution with mean β (see Sec. 6.2.2). The distribution function is

$$F(x) = \begin{cases} 1 - e^{-x/\beta} & \text{if } x \ge 0\\ 0 & \text{otherwise} \end{cases}$$

so to find F^{-1} , we set u = F(x) and solve for x to obtain

$$F^{-1}(u) = -\beta \ln \left(1 - u\right)$$

Thus, to generate the desired random variate, we first generate a $U \sim U(0, 1)$ and then let $X = -\beta \ln U$. [It is possible in this case to use U instead of 1 - U, since 1 - U and U have the same U(0, 1) distribution. This saves a subtraction.]

eg.1 令 X 具有均值为 β 的指数分布,其分布函数为:

$$F(x) = egin{cases} 1-e^{-x/eta} & ext{if } x \geq 0 \ 0 & ext{otherwise} \end{cases}$$

则为得到 F^{-1} , 设 u = F(x) 并求解 x, 以得到:

o

 $x=F^{-1}(u)=-\beta\ln(1-u)$

因此,要生成所要求的随机变量,首先生成一个 $U \sim U(0,1)$,而后令 $X = -\beta \ln U$

es430 generating random variates The inverse-transform method's validity in the continuous case was demon-strated mathematically above, but there is also a strong intuitive appeal . The density function f (x) o <u>show annotation</u>

rdance with the desired density. The inverse-transform method can also be used when X is discrete. Here the distribution function <u>show annotation</u>

hods using the idea of indexing. Generalization, Advantages, and Disadvantages of the Inverse-Transform Method Both the continuous and discrete <u>show annotation</u> ly on the dis-tribution desired. Let us now consider some general advantages and disadvantages of the inverse-transform method in both the continuous and discrete cases. One possible impedi-ment to use <u>show annotation</u>

反变换法的 缺点:

- 在连续情况下应用这种方法可能的问题是需要计算 $F^{-1}(U)$,由于无法以闭合 形式写出所要求的分布的 F^{-1} 的公式(如正态分布和伽马分布),因此有时候 简单的反变换法不一定可行;
- 对于给定分布,反变换法或许不是生成相应随机变量的最快速方法。

优点:

- 易于使用方差缩减技术。依赖于引进随机变量间的相关性,如公共随机数法和 对偶变量法 (antithetic variates)
- 易于从截断分布中产生随机变量。

or each distribution considered. <mark>Despite these possible drawbacks, there are some important advantages in using the inverse-transform method.</mark> The first is to facilitate <u>show annotation</u>

he discrete case is analogous.) ==Then an algorithm for generating an X having dis-tribution function F is as follows:== 1. Generate U, U(0, 1).2. Let* <u>show annotation</u>

8.2.2 Composition

hucany (1972).8.2.2 Composition The composition technique applies when the distribution function F from which we wish to generate can be expressed as a convex combination of other distribution functions F_1, F_2, \cdots . We would hope to be able to sample from the F_j 's more easily than from the original F.Law01323_ch08_426-487.indd Page show annotation

8.2.3 Convolution

iated methods.8.2.3 Convolution For several important distributions, the desired random variable X can be expressed as a sum of other random variables that are IID and can be generated more readily than direct generation of X. We assume that there are I <u>show annotation</u>

sses, where the distribution of X is called the m-fold convolution of the distribution of a Y_j . FIGURE 8.7Right-trapezoidal den show annotation

组合法和卷积法的区别:

- 卷积法: 假定随机变量 X 可以表示为其他随机变量的和'
- 组合法: 假设 X 的 分布函数 可以表示为其他分布函数的 (加权)和

8.3.4). See also Devroye (1988). Convolution is really an example of a more general idea, that of transforming some intermediate random variates into a fi nal variate that has the desired distribu-tion; the transformation with convolut show annotation

8.2.4 Acceptance-Rejection

前面三种方法(反变换法、组合法和卷积法)都是直接方法,即直接面对所要求的 分布或随机变量。而可用于 <u>直接方法无效或者效率不佳</u>的情形的方法:

• 舍取法

另外两种方法:

- 均匀比法
- 特性法

e with the desired density f(x). The principle of acceptance-rejection is quite general, and looking at the above algorithm in a slightly different way clarifi

es how it can be extended to generation FIGURE 8.8f (x), t(x), and r(x) show annotation

8.3 Generating Continuous Random Variates

ATING CONTINUOUS RANDOM VARIATES In this section we discuss particular algorithms for generating random variates from several commonly occurring continuous distributions; Sec. 8.4 contains a similar <u>show annotation</u>

st and greater complexity. I n deciding which algorithm to present, we have tried to choose those that are simple to describe and implement, and are reasonably effi cient as well. We also give only exact (up to machine accuracy) methods, as opposed to approximations . If speed is critically importa <u>show annotation</u>

1. Uniform 均匀分布

ions, see Secs. 6.2.2 and 6.2.3. 8.3.1 Uniform The distribution function of a U show annotation

2. Exponential

stored for use in the algorithm. <mark>8.3.2 Exponential</mark> The exponential random variable show annotation

3. *m*-Erlang

01323_pagefileschapter eight 453 <mark>8.3.3 m -Erlang</mark> As discussed in Example 8.5, if <u>show annotation</u>

rnative method when m is large. Fortunately, the m-Erlang dis-tribution is a special case of the gamma distribution (with shape parameter a equal to the integer m), so that we can use one of the methods for generating gamma

ran-dom variates here as well (see Sec. 8.3.4 for discussion of gamma generation). The precise threshold for m bey <u>show annotation</u>

4. Gamma

ies when its logarithm is taken. <mark>8.3.4 Gamma</mark> General gamma random variates show annotation

5. Weibull

h they discuss in their chap. 9. 8.3.5 Weibull The Weibull distribution functio <u>show annotation</u>

6. Normal

eibull(a, b); see Sec. 6.2.2. <mark>8.3.6 Normal</mark> First note that given X , N(<u>show annotation</u>

of much faster algorithms. It does have the advantage, however, of maintaining a one-to-one correspondence between the random numbers used and the N(0, 1) random vari-ates produced ; it may thus prove useful for m

<u>show annotation</u>

ly IID U(0, 1) random variables, there is a serious diffi culty if U_1 and U_2 are actually adjacent random numbers produced by a linear congruential generator (see Sec. 7.2), as they might be in practice. D show annotation

极坐标法 Polar method

should probably be used instead. <mark>An improvement to the Box and Muller</mark> method, which eliminates the trigono-metric calculations and was described in Marsaglia and Bray (1964), has become known as the polar method . It relies on a special propert <u>show annotation</u>

7. Lognormal

rd normal distribution function. 8.3.7 Lognormal A special property of the lognor show annotation

8. Beta

s2 5 Var(Y) 5 In a1 1 s;2m;2 b 8.3.8 Beta First note that we can obtain X9 show annotation

8.4 Generating Discrete Random Variates

E RANDOM VARIATESThis section discusses particular algorithms for generating random variates from various discrete distributions that might be useful in a simulation study . As in Sec. 8.3, we usually pre <u>show annotation</u>

e expense of greater complexity. The discrete inverse-transform method, as described in Sec. 8.2.1, can be used for any discrete distribution, whether the range of possible values is fi nite or (count-ably) infi nite . Many of the algorithms pr <u>show annotation</u>

• 别名法 (alias method)

riate from a given distribution. One other general approach should be mentioned here, which can be used for generating any discrete random variate having a fi nite range of values. This is the alias method, developed by Walker (1977) a show annotation

1. Bernoulli

- 2. Discrete Uniform
- 3. Arbitrary Discrete Distribution
- 4. Binomial
- 5. Geometric
- 6. Negative Binomial
- 7. Poisson

the desired probabilities p(i). Although the alias method is limited to discrete random variables with a fi nite range, it can be used indirectly for discrete distributions with an infi nite range, such as the geometric, negative binomial, or Poisson, by combining it with the gener <u>show annotation</u>

8.5 Generating Random Vectors, Correlated Random Variates, and Stochastic Processes

RIATES, AND STOCHASTIC PROCESSES So far in this chapter we have really considered generation of only a single random variate at a time from various univariate distributions. Applying one of these al <u>show annotation</u>

本章到目前为止,实际考虑的只是从各种 *单变量* 分布中每次产生单个随机变量。利 用随机数组重复应用这些方法的一种由所要求的分布产生一串独立同分布的随机变 量。

1. Using Conditional Distributions

tions and is left to the reader. <mark>As general as this approach may be, its practical utility is probably quite limited.</mark> Not only is specifi cation <u>show annotation</u>

2. Multivariate Normal and Multivariate Lognormal

tivariate normal distri-bution. Since Σ is symmetric and positive definite, we can factor it uniquely as $\Sigma = CC^T$ (called the Cholesky decomposition),

where the $d \times d$ matrix C is lower triangular . Algorithms to compute C can be <u>show annotation</u>

rn X 5 (eY1, eY2, ..., eYd)T. Note that μ and Σ are not the mean vector and covariance matrix of the desired multivariate lognormal random vector X, but rather are the mean and covariance matrix of the corresponding multivariate normal random vector Y. Formulas for the expected va show annotation

3. Correlated Gamma Random Variates

atesWe now come to a case where we cannot write the entire joint distribution but only specify the marginal distributions (gamma) and the correlations between the com-ponent random variables of the X vector. Indeed, there is not even agree <u>show annotation</u>

- 4. Generating from Multivariate Families
- 5. Generating Random Vectors with Arbitrarily Specified Marginal Distributions and Correlations

and CorrelationsIn Sec. 6.10.2 we noted the need to model some input random variables as a random vector with fairly arbitrary marginal distributions and correlation structure, rather than specifying and controlling their entire joint distribution as a member of some multivariate parametric family like normal, lognormal, Johnson, or Bézier . The in-dividual marginal distr

show annotation

continuous, discrete, or mixed. The only constraint is that the correlation structure between them be inter-nally consistent with the form and parameters of the marginal distributions, as dis-cussed by Whitt (1976), i.e., that the correlation structure specifi ed be feasible. The modeling fl exibility and

show annotation

orrelation structure desired. They present specifi c examples through dimension d 5 3 when the marginal distributions are uniform, exponential, and discrete uniform .Cario et al. (2002) develop a m <u>show annotation</u>

6. Generating Stochastic Processes

sesAs mentioned in Sec. 6.10.3, some applications require that we generate observa-tions of the "same" random variable as it is observed through time. For example, we might want to g <u>show annotation</u>

常见随机过程的产生:

- AR 和 ARMA 过程
- 伽马过程
- ARTA 过程
- VARTA 过程

8.6 Generating Arrival Processes

- 1. 泊松过程
- 2. 非平稳泊松分布

9. Output Data Analysis for a Single System

9.1 Introduction

ion out-put data appropriately. As a matter of fact, a common mode of operation is to make a single simulation run of somewhat arbitrary length and then to treat the resulting simulation estimates as the "true" model characteristics. Since random samples from p <u>show annotation</u>

9.2 Transient and Steady-state Behavior of a Stochastic Process

ry much like the density fYij. For fixed y and I, the probabilities $F_1(y|I), F_2(y|I), \cdots$ are just a sequence of numbers. If $F_i(y|I)$ as $i \to \infty$ for all y and for any initial conditions I, then F(y) is called the steady-state distribution of the output process $Y_1, Y_2, \cdots, Y_2, \ldots$. Strictly speak-ing, <u>show annotation</u>

Note: 稳态并不意味着在一次特定的仿真中随机变量 *Y*_{*k*+1}, *Y*_{*k*+2}, · · · 都取相同的 值,而是意味着它们都 近似地 具有相同的分布。进一步,这些随机变量之间并不独 立,但近似一个 *方差平稳* 的随机过程。

9.3 Types of Simulations with Regard to Output Analysis

S WITH REGARD TO OUTPUT ANALYSIS The options available in designing and analyzing simulation experiments depend on the type of simulation at hand, as depicted in Fig. 9.4 . Simulations may be either term <u>show annotation</u>

• 终止型仿真 terminating simulation

nt runs are IID (see Sec. 9.4). The event E often occurs at a time point when the system is "cleaned out" (see Example 9.4), at a time point beyond which no useful information is obtained (see Example 9.5), or at a time point specifi ed by management mandate (see Example 9.8). It is s <u>show annotation</u>

• 非终止型仿真 nonterminating simulation

ith the current inventory level. <mark>A nonterminating simulation is one for which there is no natural event E to specify the length of a run.</mark> This often occurs when we are d <u>show annotation</u>

ata analysis for a single system <mark>It should be mentioned that stochastic</mark> processes for most real systems do not have steady-state distributions, since the characteristics of the system change over time . For example, in a manufacturin <u>show annotation</u>

) may change from time to time. On the other hand, a simulation model (which is an abstraction of reality) may have steady-state distributions, since characteristics of the model are often assumed not to change over time. When we have new information on the characteristics of the system, we can redo our steady-state analysis. If, in Example 9.9, the man <u>show annotation</u>

e delay over a week, nC 5 E(DC). For a nonterminating simulation, suppose that the stochastic process Y1, Y2, . . . does not have a steady-state distribution, and that there is no appropriate cycle defi -nition such that the corresponding process Y1C, Y2C, . . . has a steady-state distribu-tion . This can occur, for example, i <u>show annotation</u>

probably not be well defi ned. In these cases, however, there will typically be a fi xed amount of data describing how input parameters change over time. This provides, in effect, a te <u>show annotation</u>

9.4 Statistical Analysis for Terminating Simulations

YSIS FOR TERMINATING SIMULATIONS Suppose that we make *n* independent replications of a terminating simulation, where each replication is terminated by the event *E* and is begun with the "same" initial conditions (see Sec. 9.4.3). The independe <u>show annotation</u>

Obtaining a specified precision

Obtaining a Specifi ed Precision One disadvantage of the fixed-sample-size procedure based on n replications is that the analyst has no control over the confidence-interval half-length (or the precision of X(n)); for fixed n, the half-length will depend on Var(X), the population variance of the X_j 's . In what follows we discus show annotation